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Enhancing the Inverted Topp-Leone Geometric Distribution: Properties and Applications in Computing

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Abstract: Statistics has developed a substantial interest in lifetime models, specifically within the domain of statistical inference. Practical domains such as computer science, medicine, engineering, biology science, management, and public health make extensive use of these models. Probability models find application in various domains, including game-winner prediction, team classification, winning margin evaluation, and likelihood of team victory. Recently, the model of mixed distribution has gained widespread recognition in the field of statistical data modelling. This paper aims a new two-parameter generalization of inverted Topp-Leone distribution. The new model, known as the Inverted Topp-Leone Geometric distribution, is created by mixing inverted Topp-Leone and geometric distributions. The quantile function, incomplete moments, ordinary moments, median, mode, mean residual life function, entropy, Shannon entropy, and mean deviation are some of the mathematical features of the new distribution that are obtained. Other characteristics include the mean deviation. The maximum likelihood approach is used to arrive at an estimate of the parameters of the model. Inverted (or inverse) distributions are advantageous for investigating further characteristics of the phenomenon. The behaviour of the parameter estimations is investigated by a Monte Carlo simulation. A practical computing application is provided to illustrate the new model's usefulness.

Keywords: Inverted Topp-Leone Distribution (ITLD); Geometric Distribution (GD); Maximum Likelihood Estimation (MLE).

1. Introduction

Several lifetime continuous distributions are derived to describe and predict the real-life phenomenon. However, there is always room for the development of new mod-els that are more adaptable or more suited to certain real-world issues [1-2]. Probability distribution theory has also contributed to sample theory. Different Probability models are utilized to tackle social and real-world problems. With limited support, it displayed a continuous unimodal Topp-Leone distribution. This model was employed in subse-quent statistical analyses as a substitute for the beta distribution when modelling life-time phenomena [3-5].

The hazard rate function has a bathtub shape, and the T-L distribution's density function has a J shape. When there are two separate causes of failure, mixtures of life-time distributions with the same parametrical form of distributions arise [6]. Given the significance of TL dispersion, several writers carried out related research. To investigate other aspects of the phenomenon, the inverted (or inverse) distributions are useful [7-9]. Studies in the fields of econometrics [10], engineering sciences [11], survey sampling [12], medical applications [13], and life testing challenges all make use of inverted dis-tributions [14-17]. The right tail of the inverted TL distribution is rather lengthy.

The default rate value rises and gets closer to zero in the early stages as x increases. The goal is to conclude the proposed lot's acceptance or rejection because the units' ac-tual median lifetime, exceeds the required lifetime [18-20]. When conducting a life test, it is customary to end it at a predetermined time and record the number of failures. It may also simulate bathtub and inverted J hazards, as well as hazards that are growing and diminishing simultaneously [21]. Being able to deal with a highly controllable and simple distribution of interest with an accurate closed-form CDF is another benefit [22-25]. This distribution is a very good fit for a variety of applications because of its strong requirements.

2. Materials and Methods

2.1. Derivation of New Model

Assume that "N" is a geometric random variable with a probability density function.

 $f(n,p) = p(N = n) = (1 - u) u^{n-1}, \quad n \in N , u \in (0,1)$

Let $z = Min\{z_1, z_2, z_3, ..., z_N\}$ following a random sample taken from an inverted Topp-leone distribution.

$$G(z) = 1 - \left\{ \frac{(1+2y)^{\theta}}{(1+y)^{2\theta}} \right\}; z \ge 0,$$

$$F(y)_{New} = \frac{F(y)}{1 - p(1 - F(y))}$$
(1)

The corresponding pdf is

$$f(y) = \frac{(1-p) f(y)}{[1-p \{1-F(y)\}]^2}$$

By using Eq. (1), In the ITLG distribution, the unconditional CDF is considered to be

$$F_{ITLG}(y;\theta,p) = \frac{1 - \left\{ \frac{(1+2y)^{\theta}}{(1+y)^{2\theta}} \right\}}{1 - p \left\{ \frac{(1+2y)^{\theta}}{(1+y)^{2\theta}} \right\}}; z \ge 0, \theta$$

$$0$$
(2)

For z>0, it is expressed as the probability density function (pdf) that corresponds to this statement.

$$f_{ITLG}(y;\theta,p) = \frac{2\theta(1-p)y(1+y)^{-2\theta-1}(1+2y)^{\theta-1}}{\left[1-p\left\{\frac{(1+2y)^{\theta}}{(1+y)^{2\theta}}\right\}\right]^2}$$
(3)

For the given values of θ and p, some plots of the ITLG distribution are shown in Figures 1 and 2. In Figure 2, diagrams illustrate the HR function of ITLG Distribution.



Figure 1. ITLG Distribution's PDF for Specific Parameter Values

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Figure 2. Shows the HF of the ITLG Distribution Parameter





Figure 3. Certain parameter values for SF and ITLG Distribution

3. All Statistical Properties of the Proposed Distribution of ITLG

Linear Depiction for the ITLG Delivery System. We show a linear representation of the ITLG distribution and describe how to apply it to give a practical linear representation. This is one way to provide a mixed representation of the ITLG. Due to the com-plexity of the precise solution, numerical computations of skewness and kurtosis are performed.

θ				Р		
	0.1	0.25	0.45	0.5	0.75	0.95
5	3.5579	3.4799	3.7807	3.9308	5.5042	12.5168
6	2.7683	2.7189	3.0262	3.1689	4.5945	10.7004
7	2.3464	2.3152	2.6315	2.7716	4.1269	9.7752
8	2.7683	2.7189	3.0262	3.1689	4.5945	10.7004
9	2.3464	2.3152	2.6315	2.7716	4.1269	9.7752

Table 1. Parameter Combination Skewness in the ITLG Distribution

Mode: As a result of Equation (4), the rth moment of z is calculated.

$$E(y^r) = \int_0^\infty y^r f(y) dy$$
(5)

Mean: The formula for calculating the mean of the ITLG distribution is as follows:

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$$\mu'_{1} = 2\theta(1-p)\sum_{k=0}^{\infty}\sum_{t=0}^{\infty}\phi_{k,t}\beta(3+t,2\theta(k+1)-\theta(k+1)-1)$$
(6)

Median: The median of the ITLG distribution is given as $(1 + z)^{2\theta}(1 - 0.5) + (1 + 2z)^{\theta}(P(0.5) - 1) = 0$

$$(1+z)^{2\theta}(1-0.5) + (1+2z)^{\theta}(P(0.5)-1) = 0$$
(7)

A description of the manner of the ITLG distribution is as follows: $[1 - 2\theta y^2 - 2y^2][(1 + y)^{2\theta} - p(1 + 2y)^{\theta}] = 4p\theta y^2(1 + 2y)^{\theta}$ (8)

The definition of the mgf of the ITLG distribution is as follows:

$$M_{z}(t) = E(e^{tz}) = \int_{0}^{\infty} e^{tz} f(z) dz$$
$$M_{z}(t) = 2\theta(1-p) \sum_{r=0}^{\infty} \sum_{k=0}^{\infty} \sum_{t=0}^{\infty} \phi_{k,t,r} \beta(r+t+2,2\theta(k+1)-\theta(k+1)-r)$$

Where

$$\phi_{k,t,r} = \frac{t^r}{r!} \frac{\Gamma\left(2+k\right)}{k!} p^k \begin{pmatrix} \theta(k+1) - 1 \\ t \end{pmatrix}$$

Mainly because the computation of variance is challenging and complex. Consequently, to compute the variance in Table 3, we make use of a numerical estimation approach.

Θ —		Р								
	0.1	0.25	0.45	0.5	0.75	0.95				
5	0.5252	0.5081	0.4352	0.4089	0.2346	0.0511				
6	0.3302	0.326	0.284	0.2677	0.1553	0.034				
7	0.2316	0.2321	0.2048	0.1934	0.1131	0.0248				
8	0.1741	0.1766	0.1573	0.1488	0.0876	0.0193				
9	0.1372	0.1406	0.1262	0.1195	0.0706	0.0156				

|--|

3.1. Cumulative HRF

$$H(y) = \int_{-\infty}^{x} h(u) du = -\log [1 - F(y)] = -\log(S(y))$$
$$= \log \left[1 - p \left\{ \frac{(1 + 2y)^{\theta}}{(1 + y)^{2\theta}} \right\} \right] - \log \left[(1 - p) \left\{ \frac{(1 + 2y)^{\theta}}{(1 + y)^{2\theta}} \right\} \right]$$

3.2. Mean Deviation

$$M.D(y) = \int_0^\infty |y - \mu| f(y) dz$$
$$M.D(z) = 2\mu \frac{1 - \left\{\frac{(1+2y)^\theta}{(1+z)^{2\theta}}\right\}}{1 - p\left\{\frac{(1+2y)^\theta}{(1+y)^{2\theta}}\right\}} - 4\theta(1-p) \sum_{k=0}^\infty \sum_{t=0}^\infty \phi_{k,t} \beta_{\frac{\mu}{1-\mu}}(t+3, 2\theta(k+1) - \theta(k+1) - 1)$$

3.3. Shannon Entropy

Entropy is an information metric that provides a quantified range of probable outcomes. Shannon established it in 1948. It is a measure of the average probable outcomes of a random variable [26-27].

It represents the mean of all conceivable outcomes for a random variable. Entropy reaches its maximum indicating that there isn't any uncertainty in its characterization [28].

Given an uncontrolled variable Z, a set z serving as its support, and the probability density function f(z), the Shannon's entropy of Z is

$$H_z = -\int_{-\infty}^{\infty} f(y) \log f(y) \, dz$$

Integrals' presence is one of our main requirements. Thus, the ITLG random variable Z's differential is Rényi Entropy.

3.4. Rényi Entropy

The Rényi entropy is widely used in computer science, econometrics, statistical inference, classification, and identification of statistical problems [29]. Z is given the Rényi entropy as though it adhered to the ITLG distribution.

$$= (1-p)^{\nu} (2\theta)^{\nu} \sum_{k=0}^{\infty} \sum_{t=0}^{\infty} \phi_{k,t} \beta(\nu+t+1,\nu(\theta+2)+\theta k-\nu-1)$$

Let Z be the parameter that corresponds to the observed distribution and let p, (θ) [^]z be the dimension of the parametric matrix. By [30], the sample LF is computed as follows:

For each sample size, the tests were run 10,000 times.

Tables 4, 5, and 6 display the results of the MC simulation. Calculations are made for the variance of estimated parameters, biases, mean of estimated characteristics, and mean square error [31]. These results are grounded in the predicted first-degree asymptotic theory, which states that bias and MSE decrease to zero as the sample size increases.

	Table 5. Absolute bias, variance and MSE at $unu p = 0.5, \theta = 0.5$							
Random	3	0	10	00	30	300		00
Sample								
Parameters	р	Θ	р	θ	Р	θ	р	θ
Estimate	0.4939	0.5364	0.4977	0.5109	0.4992	0.5042	0.4994	0.5019
Bias	0.006	0.0364	0.0022	0.0109	0.0007	0.0042	0.0002	0.0019
Variance	0.003	0.0226	0.0008	0.0056	0.0002	0.0018	0.0001	0.001
Mean	0.003	0.0212	0.0008	0.0057	0.0002	0.0018	0.0001	0.001
Square Error								

Table 3. Absolute Bias, Variance and MSE at *and* p = 0.5, $\theta = 0.5$

Table 4. Absolute Bias, variance and MSE $p = 0.5$, $\theta = 2$								
Random	3	0	1(00	30	300		00
Sample								
Parameters	р	θ	р	θ	р	θ	р	θ
Estimate	0.494	2.1387	0.4986	2.0395	0.4992	2.0139	0.4993	2.0095
Bias	0.0059	0.1387	0.0031	0.0395	0.0007	0.0139	0.0006	0.0095
Variance	0.003	0.3483	0.0008	0.088	0.0002	0.0289	0.0001	0.0171
Mean	0.0031	0.3675	0.0008	0.0895	0.0002	0.0291	0.0001	0.0172
Square								
Error								

	Table 5. Bias, Variance, and MSE p=0.9, =1.9							
Random	30		100		300		500	
Sample								
Parameters	р	Θ	р	θ	р	θ	р	θ
Estimate	0.8988	2.2729	0.8995	2.0031	0.8998	1.9315	0.8999	1.9179
Bias	0.0011	0.3729	0.0004	0.1031	0.0001	0.0315	0	0.0179
Variance	0.0001	1.6123	0	0.2409	0	0.0625	0	0.0347
Mean	0.0001	1.7514	0	0.2516	0	0.0635	0	0.035
Square								
Error								

4. Practical Application

If any outliers or medians are present in a data collection, boxplots are typically employed to demonstrate them. The boxplot shows that there is good agreement be-tween the data points in the overall plot of the breaking stress of carbon fibers.

3.70, 2.740, 2.730, 3.110, 3.270, 2.870, 4.420, 2.410, 3.190, 3.280, 3.090, 1.870, 3.750, 2.430, 2.950, 2.960, 2.30, 2.670, 3.390, 2.81000, 4.2000, 3.310, 3.310, 2.850, 3.150, 2.350, 2.550, 2.810, 2.770, 2.170, 1.410, 3.680, 2.970, 2.760, 4.910, 3.680, 3.190, 1.5700, 0.8010, 1.590, 2.000, 1.220, 2.170, 1.170, 5.080, 3.510, 2.170, 1.690, 1.8400, 0.390, 3.680, 1.610, 2.790, 4.70, 1.570, 1.080, 2.030000, 1.890, 2.880, 2.820, 2.50, 3.60, 1.4700, 3.110, 3.220, 1.690, 3.150, 4.90, 2.970, 3.390, 2.930, 3.220, 3.330, 2.550, 2.560, 3.56000, 2.590, 2.380, 2.830, 1.920, 1.360, 0.980, 1.840, 1.590, 5.560, 1.7300, 1.120, 1.710, 2.480, 1.180, 1.250, 4.380, 2.480, 0.85000, 2.030, 1.80, 1.610, 2.120, 2.050, 3.650



Figure 4. Estimated Densities of Data

Additionally, Figures 4 and 5 provide plots of fitted densities and empirical curve fits versus the MOIW, IPL, IK, and GIL distributions. These figures demonstrate that the ITLG distribution offers the most accurate representation of the actual data.



Figure 5. Empirical CDF Curves for MOE Distribution.

Table 7: MLEs of the Dataset					
Model	Estimates o	f (Standard Deviations)			
ITLG	$\theta = 71.4130 (3.15030)$	p = 0.66520 (0.15340)	-		

			$\theta =$
MOIW	$\alpha = 5.74940 \ (0.79520)$	$\beta = 2.23700 \ (0.000800)$	0.08970
			(0.01160)
			$\theta =$
IPL	$\alpha = 1.98740 \ (0.15660)$	$\beta = 0.23170 \ (0.00870)$	0.52860
			(0.32670)
IK	$\alpha = 2.75190 \ (0.21130)$	$\beta = 2.92270 \ (0.11190)$	-
GIL	$\alpha = 0.38420 \ (0.03840)$	$\theta = 1.26540 \ (0.24170)$	-

Table 8. Goodness of Fit Quantities of Dataset								
Model	ITLG	MOIW	IPL	IK	GIL			
-2logL	285.24	288.23	287.48	286.28	391.12			
AIC	291.24	296.23	293.48	290.28	395.12			
BIC	299.06	306.65	301.29	298.49	400.33			
HQIC	294.41	300.44	296.64	294.91	397.22			
K-S	0.0539	0.0966	0.088	0.0823	0.2521			
MVS	0.0631	0.0841	0.0783	0.0724	0.1934			
AD	0.413	0.612	0.598	0.561	0.92			
p-value	0.9326	0.7981	0.8124	0.8329	0.3691			

Based on an evaluation of all the models mentioned above, it is found that, for the given real-world data set, the ITLG distribution outperforms all other inverted distribution

5. Conclusion

The aforementioned table displays the values of ML estimates along with the six information requirements. The ITLG distribution provides the best fit out of MOIW, IPL, IK, and GIL. Compared to the distributions of MOIW, IPL, IK, and GIL, ITLG exhibits lower values for the aforementioned criterion. Final Thoughts on an Application. We illustrate the application of the inverted TLG distribution using real-world data and contrast it with MOIW, IPL, IK, and GIL models. The features are predicted using the MLE technique. The performance of MLE is assessed by simulation experiments. The comparison is made using a range of goodness of fit criteria and graphical analysis.

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